

2023 South Carolina Adult Tobacco Survey Small Area Estimates

Tobacco Prevention and Control Section

South Carolina Department of Public Health



**Prepared by:
RTI International
Research Triangle Park, NC**

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1. Overview

The 2023 South Carolina Adult Tobacco Survey (SC—ATS) was designed to produce precise state-level estimates but does not have adequate sample size to make direct estimates to small domains with adequate precision. Direct estimates only use data related to the domain in question and the study design information (e.g., sampling weights and stratification). We apply a modeling approach to make indirect estimates to small domains. The statistical model leverages a model to make “indirect” estimates for domains that lack adequate sample size for direct estimation.

The small areas described in this report include the 46 individual South Carolina counties and 32 domains composed of the cross-classification of sex, age category (18-24, 25-44, 45-64, 65 and older), and race/ethnicity (White NH, Black NH, Hispanic, other NH).

For each small area, estimates and 95% confidence intervals are created for the following tobacco related outcomes:

- Ever cigarette smoking
- Current cigarette smoking
- Ever e-cigarette use
- Current e-cigarette use
- Secondhand smoke exposure at work
- Secondhand smoke exposure in public

The modeling approach we applied is similar to the approach described in Srebotnjak et al. (2010) with differences described in Appendix A. Appendix A is separated from the main report because most readers will not have the technical background to assess the information discussed.

2. Methodology

2.1 County-Level Estimates

For each of the six outcomes described in Section 1, our method of calculating county-level estimates include the following steps:

- Weighted logistic models were fit to the 2023 SC—ATS data, one model for each outcome. The independent variables in the models were: region, age category, sex, and educational attainment. The dependent variable is the outcome.
- For each South Carolina County, a dataset with 60 records was created from the cross-classification of the following variables: sex (two categories), age (five categories), race/ethnicity (two categories), educational attainment (three categories). These data are referred to as synthetic data. They are called synthetic data because each data point is created in statistical analysis software, in contrast to data collected from survey respondents.
- For each synthetic datapoint an estimate of the probability of observing that outcome was calculated using the fitted model parameters.
- For each county, the synthetic data is calibrated to the county’s marginal distribution totals of sex, age category, race/ethnicity category, and educational attainment category, obtained from U.S. Census Bureau data.
- The weighted synthetic data is used to estimate the prevalence of each outcome for every county.
- The precision of the estimates (i.e., 95% confidence intervals) is calculated using a statistical technique called bootstrap variance estimation described in Appendix B.

2.2 Sex by Age Category by Race/Ethnicity Estimates

For each of the six outcomes described in Section 1, we applied a method to calculate estimates for the 32 domains composed of the cross-classification of sex, age category, and race/ethnicity that includes the following steps:

- Weighted logistic models were fit to the 2023 SC—ATS data, one model for each outcome. The independent variables in the models were: region, age category, sex, and educational attainment. The dependent variable is the outcome.
- A dataset with 768 records was created from the cross-classification of the following variables: region (4 categories), sex (2 categories), age (6 categories), race/ethnicity (4 categories), educational attainment (4 categories). These data are referred to as synthetic data.
- For each synthetic datapoint an estimate of the probability of observing that outcome was calculated using the fitted model parameters.

- The synthetic data is calibrated to the state’s marginal distribution totals of sex, age category, race/ethnicity category, and educational attainment category, obtained from U.S. Census Bureau data.
- The weighted synthetic data is used to estimate the prevalence of each outcome for each of the 32 domains composed of the cross-classification of sex, age category (18-24, 25-44, 45-64, 65 and older), and race/ethnicity (White NH, Black NH, Hispanic, other NH).
- The precision of the estimates (i.e., 95% confidence intervals) is calculated using a statistical technique called bootstrap variance estimation described in Appendix B.

3. Data Used in the Analysis

In 2023, the SC—ATS was conducted using an address-based sample (ABS) survey. The sample was allocated to maximize the precision of the state-level estimates; there was no geographic stratification and no oversampling of any geographies. The sample was a simple random sample of addresses on the USPS Delivery Sequency File. In expectation, the number of respondents in each county is proportional to the number of addresses in each county. There were 1,370 respondents across the 46 counties.

Exhibits 3.1 and 3.2 display the number of 2023 SC—ATS respondents for each of the domains to which we are making small area estimates. All of the domains, to which we are making small area estimates, have less than 200 respondents. If there were more respondents in some of the domains than it would be useful to make direct estimates to compare with the indirect estimates. Comparing the direct and indirect estimates, for the domains with adequate sample sizes, is a method of evaluating the correctness of the model and ultimately the correctness of the indirect estimates. However, with only one year of data, we do not have enough sample in any of the domains to conduct this evaluation. After two years of data collection, we will have adequate sample sizes in some of the domains to make this evaluation.

Table 1. Number of SC--ATS Respondents by County

Displays the number of SC--ATS respondents by county. The counties are ordered by the number of respondents.

| County name | Adult population | 2023 SC--ATS respondents |
|--------------------|-------------------------|---------------------------------|
| Richland | 327,281 | 148 |
| Greenville | 425,058 | 125 |
| Charleston | 338,978 | 110 |
| Horry | 316,837 | 100 |
| York | 226,190 | 81 |
| Spartanburg | 266,429 | 69 |
| Lexington | 237,270 | 63 |
| Berkeley | 188,229 | 62 |
| Dorchester | 127,435 | 54 |
| Anderson | 163,262 | 48 |
| Aiken | 137,853 | 47 |
| Beaufort | 161,504 | 43 |
| Sumter | 79,752 | 36 |
| Florence | 105,110 | 30 |
| Orangeburg | 64,852 | 28 |
| Oconee | 65,322 | 26 |
| Greenwood | 53,589 | 25 |
| Georgetown | 54,164 | 20 |
| Laurens | 53,336 | 20 |
| Pickens | 106,087 | 20 |
| Cherokee | 43,435 | 19 |
| Chesterfield | 34,273 | 16 |
| Lancaster | 82,828 | 15 |
| Kershaw | 52,526 | 14 |
| Darlington | 48,668 | 13 |
| Clarendon | 25,350 | 11 |
| Newberry | 29,942 | 11 |
| Hampton | 14,478 | 10 |
| Marion | 22,247 | 10 |
| Colleton | 30,064 | 9 |
| Edgefield | 22,472 | 9 |
| Williamsburg | 24,411 | 9 |
| Barnwell | 15,650 | 7 |
| Jasper | 26,393 | 7 |
| Marlboro | 20,954 | 7 |
| Lee | 13,104 | 6 |
| Saluda | 14,902 | 6 |
| Union | 21,303 | 6 |
| Abbeville | 19,558 | 5 |
| Fairfield | 16,858 | 5 |
| McCormick | 8,777 | 5 |
| Bamberg | 10,408 | 4 |
| Chester | 25,028 | 4 |
| Dillon | 20,896 | 4 |
| Calhoun | 11,599 | 2 |
| Allendale | 6,255 | 1 |
| Over all counties | 4,160,917 | 1,370 |

Table 2. Number of SC--ATS Respondents by County

Displays the number of SC—ATS respondents by sex, by age category, and by race/ethnicity category.

| Sex | Age category | Race/ethnicity | Population | 2023 SC-ATS Respondents |
|--------------|---------------------|-----------------------|-------------------|--------------------------------|
| Female | 18 to 24 | White NH | 134,831 | 23 |
| | | Black NH | 71,108 | 15 |
| | | Hispanic | 19,693 | 3 |
| | | Other NH | 13,310 | 2 |
| | 25 to 44 | White NH | 405,613 | 156 |
| | | Black NH | 192,269 | 81 |
| | | Hispanic | 46,084 | 19 |
| | | Other NH | 32,104 | 9 |
| | 45 to 64 | White NH | 456,077 | 196 |
| | | Black NH | 179,003 | 99 |
| | | Hispanic | 30,218 | 13 |
| | | Other NH | 22,462 | 9 |
| 65 and older | White NH | 415,182 | 175 | |
| | Black NH | 120,040 | 47 | |
| | Hispanic | 10,597 | 3 | |
| | Other NH | 12,566 | 2 | |
| Male | 18 to 24 | White NH | 144,004 | 13 |
| | | Black NH | 70,396 | 7 |
| | | Hispanic | 22,758 | 4 |
| | | Other NH | 14,330 | 1 |
| | 25 to 44 | White NH | 400,632 | 83 |
| | | Black NH | 172,551 | 29 |
| | | Hispanic | 54,428 | 5 |
| | | Other NH | 28,832 | 11 |
| | 45 to 64 | White NH | 440,768 | 128 |
| | | Black NH | 147,417 | 31 |
| | | Hispanic | 34,796 | 7 |
| | | Other NH | 19,219 | 8 |
| 65 and older | White NH | 348,239 | 154 | |
| | Black NH | 82,530 | 29 | |
| | Hispanic | 9,124 | 4 | |
| | Other NH | 9,731 | 4 | |
| Total | | 4,160,912 | 1,370 | |

Table 3. Tobacco Outcome Estimates for South Carolina

Displays the state-level prevalence for the 6 tobacco outcome estimated in this report.

| Tobacco outcome | n | % | 95% confidence interval |
|----------------------------|----------|----------|--------------------------------|
| Ever smoker | 1,370 | 68.9 | (65.5, 72.4) |
| Current smoker | 1,370 | 11.8 | (9.5, 14.2) |
| Ever e-cigarette user | 1,370 | 20.4 | (17.2, 23.6) |
| Current e-cigarette user | 1,370 | 7.5 | (5.4, 9.7) |
| Secondhand smoke at home | 1,370 | 8.3 | (6.3, 10.3) |
| Secondhand smoke in public | 1,370 | 28.6 | (25.4, 31.9) |

4. County-Level Estimates

Table 4. Ever Smoking and Current Smoking Estimates (%) and 95% Confidence Intervals for South Carolina Counties

| FIPS | County name | Population | Ever smoking | | Current smoking | |
|------|--------------|------------|--------------|--------------|-----------------|-------------|
| | | | Estimate (%) | 95% CI | Estimate (%) | 95% CI |
| 1 | Abbeville | 19,558 | 71.0 | (65.4, 76.7) | 15.8 | (10.4,21.2) |
| 3 | Aiken | 137,853 | 72.9 | (66.1, 79.7) | 12.7 | (7.3,18.2) |
| 5 | Allendale | 6,255 | 67.0 | (59.2, 74.8) | 9.3 | (4.5,14.1) |
| 7 | Anderson | 163,262 | 71.0 | (65.3, 76.7) | 15.5 | (10.4,20.5) |
| 9 | Bamberg | 10,408 | 68.1 | (60.9, 75.3) | 8.0 | (4.1,11.9) |
| 11 | Barnwell | 15,650 | 69.9 | (63.1, 76.6) | 9.0 | (4.9,13.2) |
| 13 | Beaufort | 161,504 | 72.0 | (65.5, 78.5) | 6.7 | (3.5,9.8) |
| 15 | Berkeley | 188,229 | 70.0 | (63.4, 76.7) | 8.3 | (4.5,12.2) |
| 17 | Calhoun | 11,599 | 70.9 | (64.3, 77.6) | 8.4 | (4.5,12.3) |
| 19 | Charleston | 338,978 | 69.7 | (62.9, 76.4) | 7.0 | (3.7,10.2) |
| 21 | Cherokee | 43,435 | 70.6 | (64.8, 76.4) | 16.6 | (11.0,22.2) |
| 23 | Chester | 25,028 | 72.7 | (65.7, 79.6) | 14.3 | (8.3,20.3) |
| 25 | Chesterfield | 34,273 | 66.4 | (59.2, 73.6) | 13.8 | (8.4,19.3) |
| 27 | Clarendon | 25,350 | 65.6 | (58.7, 72.5) | 12.6 | (7.9,17.4) |
| 29 | Colleton | 30,064 | 70.9 | (64.3, 77.6) | 9.3 | (5.0,13.6) |
| 31 | Darlington | 48,668 | 65.1 | (58.2, 71.9) | 12.3 | (7.7,16.8) |
| 33 | Dillon | 20,896 | 64.2 | (57.0, 71.4) | 13.4 | (8.4,18.4) |
| 35 | Dorchester | 127,435 | 70.1 | (63.5, 76.7) | 8.2 | (4.5,11.9) |
| 37 | Edgefield | 22,472 | 72.5 | (65.4, 79.5) | 14.1 | (8.1,20.1) |
| 39 | Fairfield | 16,858 | 71.1 | (63.7, 78.4) | 13.4 | (7.6,19.2) |
| 41 | Florence | 105,110 | 64.1 | (57.3, 70.9) | 11.9 | (7.6,16.1) |
| 43 | Georgetown | 54,164 | 68.4 | (62.0, 74.7) | 10.8 | (6.8,14.8) |
| 45 | Greenville | 425,058 | 68.8 | (63.1, 74.5) | 13.6 | (9.2,18.1) |
| 47 | Greenwood | 53,589 | 68.7 | (63.0, 74.5) | 14.6 | (9.6,19.6) |
| 49 | Hampton | 14,478 | 68.9 | (61.8, 76.0) | 9.5 | (5.0,14.0) |
| 51 | Horry | 316,837 | 69.5 | (63.1, 76.0) | 11.6 | (7.0,16.2) |
| 53 | Jasper | 26,393 | 70.1 | (63.3, 77.0) | 8.8 | (4.7,12.9) |
| 55 | Kershaw | 52,526 | 73.3 | (66.5, 80.0) | 13.6 | (7.9,19.4) |
| 57 | Lancaster | 82,828 | 73.7 | (67.1, 80.4) | 12.9 | (7.3,18.4) |
| 59 | Laurens | 53,336 | 70.3 | (64.6, 76.0) | 16.2 | (10.7,21.6) |
| 61 | Lee | 13,104 | 62.6 | (55.2, 70.1) | 12.7 | (8.0,17.4) |
| 63 | Lexington | 237,270 | 73.2 | (66.4, 80.0) | 12.5 | (7.0,18.0) |
| 65 | McCormick | 8,777 | 71.9 | (66.1, 77.7) | 14.3 | (8.8,19.9) |
| 67 | Marion | 22,247 | 63.7 | (56.5, 71.0) | 12.9 | (8.1,17.7) |
| 69 | Marlboro | 20,954 | 63.7 | (56.2, 71.2) | 13.7 | (8.5,18.9) |
| 71 | Newberry | 29,942 | 72.8 | (65.9, 79.7) | 13.7 | (7.9,19.4) |
| 73 | Oconee | 65,322 | 72.7 | (67.1, 78.4) | 14.7 | (9.9,19.5) |
| 75 | Orangeburg | 64,852 | 67.4 | (60.1, 74.8) | 7.9 | (4.0,11.8) |
| 77 | Pickens | 106,087 | 70.4 | (64.5, 76.2) | 14.5 | (9.8,19.3) |
| 79 | Richland | 327,281 | 67.1 | (59.0, 75.3) | 10.8 | (5.8,15.8) |
| 81 | Saluda | 14,902 | 73.3 | (66.4, 80.1) | 14.2 | (8.2,20.2) |
| 83 | Spartanburg | 266,429 | 69.2 | (63.5, 74.9) | 15.0 | (10.0,20.0) |
| 85 | Sumter | 79,752 | 63.2 | (56.3, 70.1) | 11.8 | (7.7,16.0) |
| 87 | Union | 21,303 | 70.6 | (64.7, 76.4) | 16.7 | (10.9,22.5) |
| 89 | Williamsburg | 24,411 | 63.0 | (55.6, 70.4) | 12.8 | (8.0,17.5) |
| 91 | York | 226,190 | 72.6 | (65.7, 79.5) | 12.1 | (6.8,17.5) |

Figure 4.1. Map of Ever Smoking Estimates for South Carolina Counties

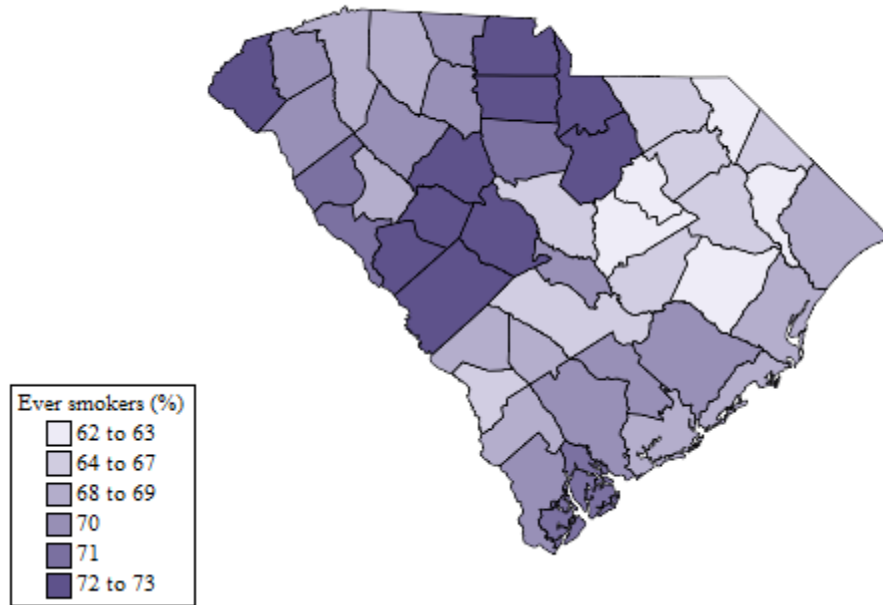


Figure 4.2. Map of Current Smoking Estimates for South Carolina Counties

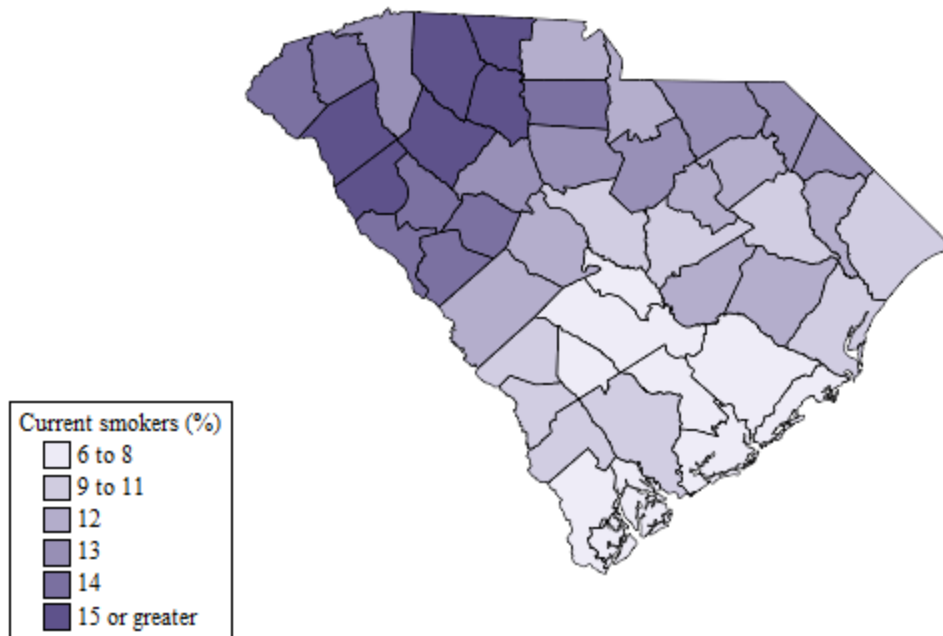


Table 5. Ever E-cigarette Use and Current E-cigarette Use Estimates (%) and 95% Confidence Intervals for South Carolina Counties

| FIPS | County name | Population | Ever e-cigarette use | | Current e-cigarette use | |
|------|--------------|------------|----------------------|--------------|-------------------------|------------|
| | | | Estimate (%) | 95% CI | Estimate (%) | 95% CI |
| 1 | Abbeville | 19,558 | 24.1 | (18.9, 29.3) | 7.2 | (3.9,10.5) |
| 3 | Aiken | 137,853 | 14.7 | (8.7, 20.8) | 7.8 | (2.9,12.6) |
| 5 | Allendale | 6,255 | 17.7 | (10.6, 24.8) | 8.8 | (3.4,14.3) |
| 7 | Anderson | 163,262 | 26.2 | (20.7, 31.7) | 7.8 | (4.2,11.3) |
| 9 | Bamberg | 10,408 | 16.9 | (10.9, 23.0) | 8.7 | (3.7,13.6) |
| 11 | Barnwell | 15,650 | 19.8 | (13.1, 26.5) | 10.0 | (4.6,15.4) |
| 13 | Beaufort | 161,504 | 17.3 | (12.0, 22.7) | 7.8 | (3.9,11.7) |
| 15 | Berkeley | 188,229 | 22.0 | (15.1, 28.8) | 10.9 | (5.3,16.6) |
| 17 | Calhoun | 11,599 | 17.8 | (11.8, 23.9) | 8.6 | (4.0,13.1) |
| 19 | Charleston | 338,978 | 20.3 | (14.2, 26.5) | 9.5 | (4.8,14.2) |
| 21 | Cherokee | 43,435 | 26.9 | (21.1, 32.6) | 8.2 | (4.5,11.9) |
| 23 | Chester | 25,028 | 15.1 | (8.8, 21.3) | 8.2 | (3.1,13.3) |
| 25 | Chesterfield | 34,273 | 20.9 | (14.8, 27.0) | 5.7 | (2.0,9.3) |
| 27 | Clarendon | 25,350 | 18.1 | (12.8, 23.3) | 4.9 | (1.9,7.9) |
| 29 | Colleton | 30,064 | 20.1 | (13.3, 26.8) | 9.8 | (4.6,15.0) |
| 31 | Darlington | 48,668 | 19.6 | (14.0, 25.2) | 5.4 | (2.1,8.6) |
| 33 | Dillon | 20,896 | 19.9 | (14.0, 25.8) | 5.6 | (2.2,9.1) |
| 35 | Dorchester | 127,435 | 21.4 | (14.7, 28.0) | 10.5 | (5.1,15.9) |
| 37 | Edgefield | 22,472 | 15.3 | (8.9, 21.7) | 8.0 | (3.0,13.0) |
| 39 | Fairfield | 16,858 | 12.2 | (6.8, 17.5) | 6.5 | (2.4,10.7) |
| 41 | Florence | 105,110 | 19.5 | (13.9, 25.0) | 5.3 | (2.1,8.5) |
| 43 | Georgetown | 54,164 | 16.4 | (11.8, 21.0) | 4.1 | (1.6,6.6) |
| 45 | Greenville | 425,058 | 25.1 | (19.9, 30.3) | 7.3 | (4.0,10.6) |
| 47 | Greenwood | 53,589 | 23.9 | (18.8, 29.0) | 7.2 | (3.9,10.5) |
| 49 | Hampton | 14,478 | 19.6 | (12.5, 26.6) | 9.7 | (4.2,15.1) |
| 51 | Horry | 316,837 | 18.9 | (13.6, 24.2) | 4.8 | (1.8,7.9) |
| 53 | Jasper | 26,393 | 18.4 | (12.1, 24.7) | 8.9 | (4.1,13.7) |
| 55 | Kershaw | 52,526 | 15.4 | (9.1, 21.7) | 8.2 | (3.1,13.2) |
| 57 | Lancaster | 82,828 | 14.8 | (8.7, 20.9) | 7.6 | (2.9,12.3) |
| 59 | Laurens | 53,336 | 25.3 | (19.9, 30.8) | 7.6 | (4.2,11.1) |
| 61 | Lee | 13,104 | 18.1 | (12.5, 23.6) | 4.9 | (1.9,8.0) |
| 63 | Lexington | 237,270 | 15.8 | (9.4, 22.3) | 8.3 | (3.1,13.4) |
| 65 | McCormick | 8,777 | 17.5 | (13.4, 21.6) | 4.5 | (2.4,6.7) |
| 67 | Marion | 22,247 | 17.9 | (12.5, 23.4) | 4.9 | (1.9,8.0) |
| 69 | Marlboro | 20,954 | 19.5 | (13.6, 25.4) | 5.3 | (2.0,8.7) |
| 71 | Newberry | 29,942 | 14.9 | (8.7, 21.1) | 7.9 | (3.0,12.9) |
| 73 | Oconee | 65,322 | 24.0 | (18.9, 29.1) | 6.7 | (3.6,9.8) |
| 75 | Orangeburg | 64,852 | 17.1 | (10.9, 23.3) | 8.8 | (3.7,14.0) |
| 77 | Pickens | 106,087 | 28.5 | (22.7, 34.4) | 8.8 | (4.8,12.8) |
| 79 | Richland | 327,281 | 14.3 | (8.1, 20.5) | 8.1 | (2.9,13.3) |
| 81 | Saluda | 14,902 | 14.7 | (8.5, 20.9) | 7.6 | (2.8,12.5) |
| 83 | Spartanburg | 266,429 | 26.1 | (20.6, 31.5) | 7.9 | (4.3,11.4) |
| 85 | Sumter | 79,752 | 19.5 | (13.9, 25.0) | 5.6 | (2.2,8.9) |
| 87 | Union | 21,303 | 24.3 | (18.9, 29.6) | 7.3 | (3.9,10.7) |
| 89 | Williamsburg | 24,411 | 17.3 | (11.9, 22.8) | 4.7 | (1.8,7.7) |
| 91 | York | 226,190 | 15.5 | (9.2, 21.9) | 8.1 | (3.1,13.1) |

Figure 5.1. Map of Ever E-cigarette Use Estimates for South Carolina Counties

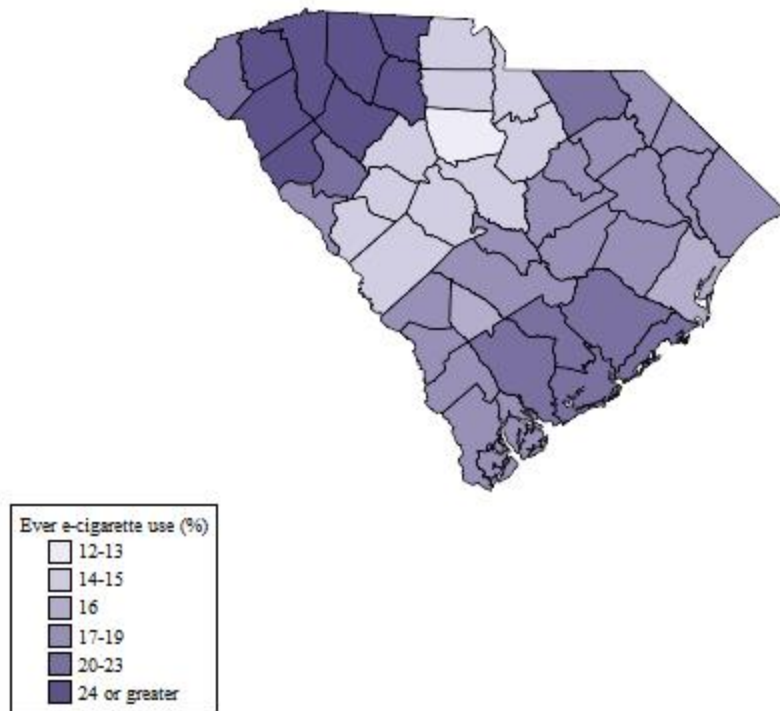


Figure 5.2. Map of Current E-cigarette Use Estimates for South Carolina Counties

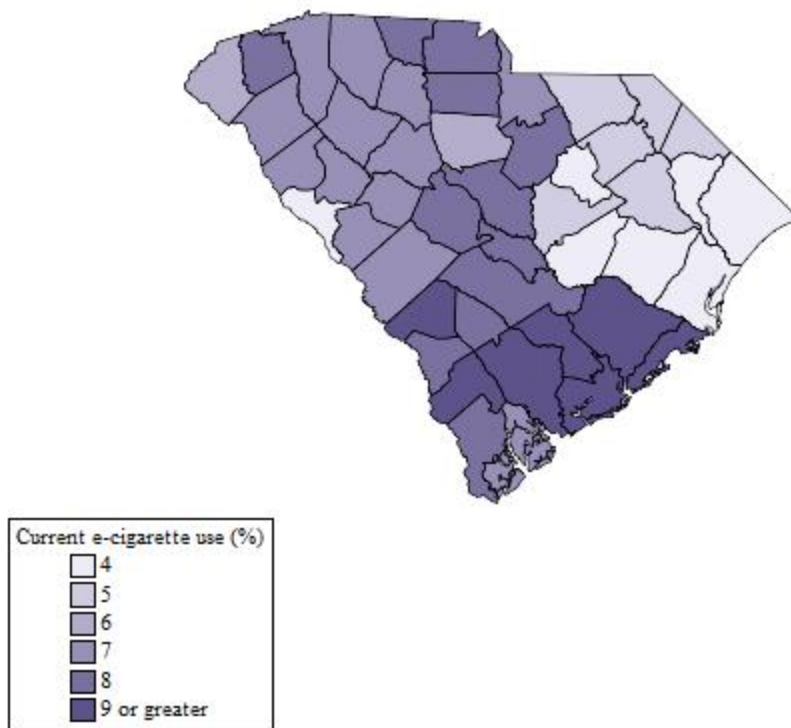


Table 6. Secondhand Smoke at Work and in Public Estimates (%) and 95% Confidence Intervals for South Carolina Counties 2023

| FIPS | County name | Population | Secondhand smoke at work | | Secondhand smoke in public | |
|------|--------------|------------|--------------------------|-------------|----------------------------|--------------|
| | | | Estimate (%) | 95% CI | Estimate (%) | 95% CI |
| 1 | Abbeville | 19,558 | 6.9 | (3.8, 10.0) | 30.5 | (23.8, 37.2) |
| 3 | Aiken | 137,853 | 12.4 | (6.4, 18.5) | 30.4 | (23.2, 37.7) |
| 5 | Allendale | 6,255 | 8.9 | (4.2, 13.5) | 27.3 | (19.9, 34.8) |
| 7 | Anderson | 163,262 | 7.4 | (4.1, 10.8) | 32.5 | (25.8, 39.3) |
| 9 | Bamberg | 10,408 | 8.3 | (4.2, 12.3) | 27.9 | (21.2, 34.7) |
| 11 | Barnwell | 15,650 | 9.5 | (5.0, 13.9) | 29.6 | (22.7, 36.6) |
| 13 | Beaufort | 161,504 | 8.0 | (4.3, 11.7) | 31.0 | (24.5, 37.5) |
| 15 | Berkeley | 188,229 | 10.3 | (5.7, 14.8) | 32.5 | (25.5, 39.5) |
| 17 | Calhoun | 11,599 | 8.7 | (4.6, 12.7) | 29.7 | (23.1, 36.4) |
| 19 | Charleston | 338,978 | 9.5 | (5.3, 13.6) | 33.5 | (26.6, 40.4) |
| 21 | Cherokee | 43,435 | 7.6 | (4.2, 11.1) | 31.8 | (24.9, 38.8) |
| 23 | Chester | 25,028 | 12.9 | (6.5, 19.3) | 29.4 | (22.4, 36.5) |
| 25 | Chesterfield | 34,273 | 6.6 | (2.5, 10.7) | 19.6 | (13.7, 25.5) |
| 27 | Clarendon | 25,350 | 5.7 | (2.3, 9.1) | 18.4 | (13.1, 23.7) |
| 29 | Colleton | 30,064 | 9.6 | (5.1, 14.0) | 30.3 | (23.3, 37.3) |
| 31 | Darlington | 48,668 | 6.2 | (2.6, 9.8) | 19.5 | (14.1, 25.0) |
| 33 | Dillon | 20,896 | 6.3 | (2.5, 10.1) | 18.9 | (13.4, 24.4) |
| 35 | Dorchester | 127,435 | 10.1 | (5.6, 14.6) | 32.8 | (25.8, 39.8) |
| 37 | Edgefield | 22,472 | 13.0 | (6.5, 19.6) | 29.9 | (22.7, 37.1) |
| 39 | Fairfield | 16,858 | 11.1 | (5.2, 17.0) | 27.5 | (20.6, 34.5) |
| 41 | Florence | 105,110 | 6.2 | (2.7, 9.7) | 19.9 | (14.5, 25.3) |
| 43 | Georgetown | 54,164 | 5.1 | (2.2, 8.1) | 19.2 | (13.9, 24.5) |
| 45 | Greenville | 425,058 | 7.1 | (3.8, 10.4) | 33.1 | (26.5, 39.6) |
| 47 | Greenwood | 53,589 | 6.8 | (3.7, 10.0) | 31.0 | (24.4, 37.5) |
| 49 | Hampton | 14,478 | 9.6 | (4.9, 14.2) | 29.3 | (22.0, 36.5) |
| 51 | Horry | 316,837 | 5.8 | (2.4, 9.2) | 20.3 | (14.5, 26.2) |
| 53 | Jasper | 26,393 | 8.7 | (4.6, 12.9) | 29.2 | (22.4, 36.0) |
| 55 | Kershaw | 52,526 | 13.0 | (6.7, 19.4) | 30.6 | (23.4, 37.9) |
| 57 | Lancaster | 82,828 | 12.7 | (6.5, 18.9) | 31.1 | (23.8, 38.3) |
| 59 | Laurens | 53,336 | 7.2 | (4.0, 10.5) | 31.3 | (24.5, 38.1) |
| 61 | Lee | 13,104 | 5.8 | (2.3, 9.3) | 18.2 | (13.0, 23.3) |
| 63 | Lexington | 237,270 | 13.4 | (6.9, 19.9) | 32.3 | (24.7, 39.9) |
| 65 | McCormick | 8,777 | 5.1 | (2.7, 7.4) | 27.6 | (21.3, 33.9) |
| 67 | Marion | 22,247 | 5.9 | (2.4, 9.3) | 18.3 | (13.0, 23.5) |
| 69 | Marlboro | 20,954 | 6.2 | (2.4, 10.1) | 18.4 | (13.0, 23.9) |
| 71 | Newberry | 29,942 | 12.6 | (6.4, 18.9) | 29.6 | (22.5, 36.7) |
| 73 | Oconee | 65,322 | 6.7 | (3.7, 9.7) | 32.1 | (25.5, 38.7) |
| 75 | Orangeburg | 64,852 | 8.4 | (4.3, 12.5) | 28.3 | (21.5, 35.1) |
| 77 | Pickens | 106,087 | 7.6 | (4.1, 11.2) | 32.7 | (25.9, 39.5) |
| 79 | Richland | 327,281 | 12.4 | (5.6, 19.2) | 30.1 | (22.4, 37.8) |
| 81 | Saluda | 14,902 | 12.7 | (6.4, 19.0) | 29.7 | (22.6, 36.9) |
| 83 | Spartanburg | 266,429 | 7.4 | (4.0, 10.8) | 32.3 | (25.6, 39.0) |
| 85 | Sumter | 79,752 | 6.1 | (2.6, 9.5) | 19.0 | (13.9, 24.2) |
| 87 | Union | 21,303 | 7.0 | (3.9, 10.2) | 30.3 | (23.4, 37.2) |
| 89 | Williamsburg | 24,411 | 5.7 | (2.3, 9.2) | 17.8 | (12.7, 22.9) |
| 91 | York | 226,190 | 13.4 | (6.8, 20.1) | 32.7 | (25.1, 40.4) |

Figure 6.1. Map of Percent Secondhand Smoke at Work Estimates for South Carolina Counties

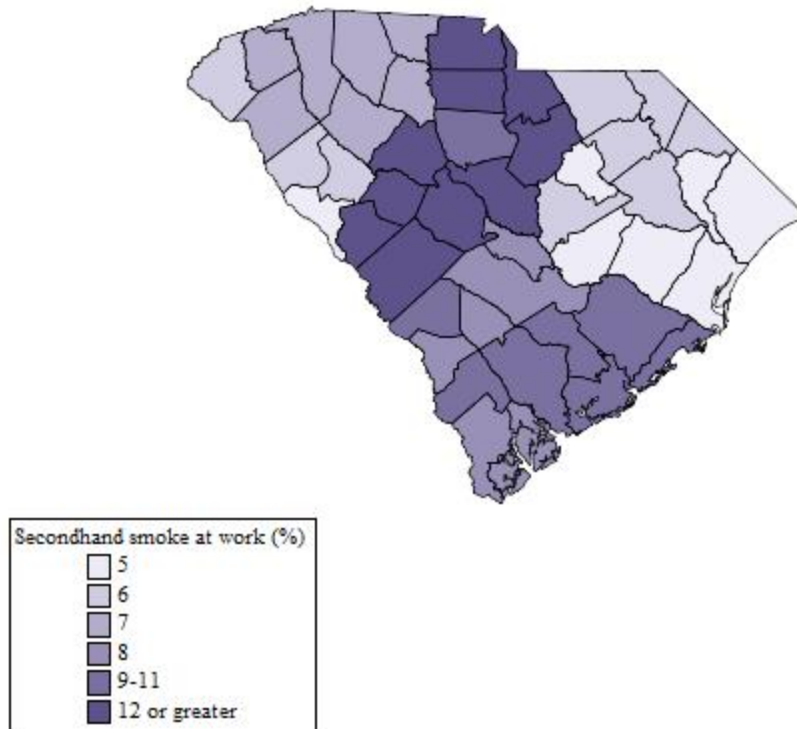
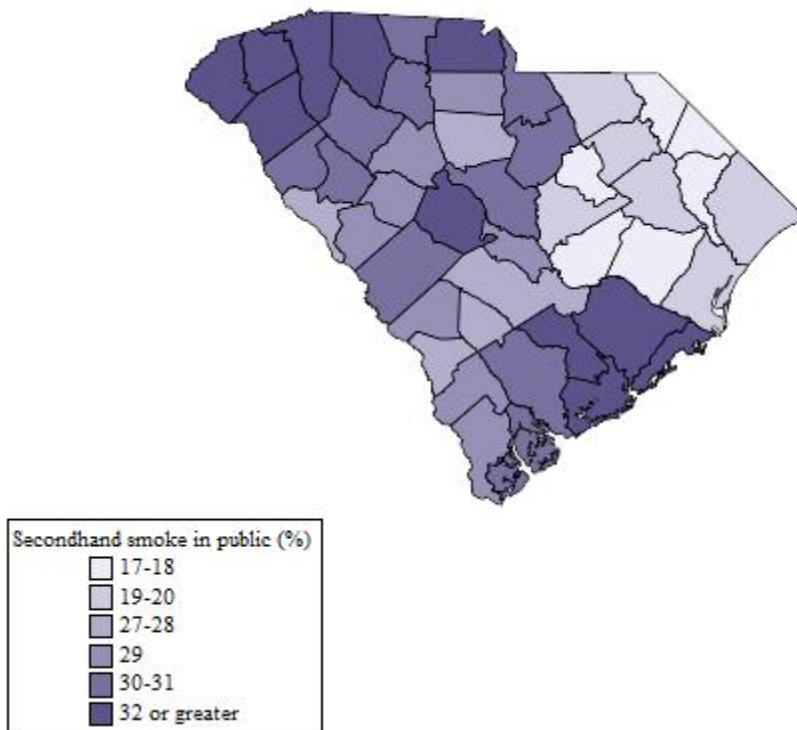


Figure 6.2. Map of Percent Secondhand Smoke in Public Estimates for South Carolina Counties



5. Tables of Sex by Age Category by Race/Ethnicity Estimates

Table 7. Ever Cigarette Use and Current Cigarette Use Estimates and 95% Confidence Intervals for Sex by Age Category by Race/Ethnicity

| Sex | Domain | | Ever cigarette use | | Current cigarette use | |
|--------|------------------------|----------------|--------------------|---------------------|-----------------------|---------------------|
| | Age category | Race/ethnicity | Estimate (%) | Confidence interval | Estimate (%) | Confidence interval |
| Female | 18–24 years old | White NH | 50.1 | (34.6, 65.7) | 7.3 | (0.0, 14.8) |
| | | Black NH | 34.1 | (19.6, 48.6) | 5.4 | (0.0, 11.0) |
| | | Hispanic | 41.6 | (19.6, 63.6) | 10.7 | (0.0, 24.7) |
| | | Other NH | 41.0 | (19.1, 62.8) | 7.7 | (0.0, 21.2) |
| | 25–44 years old | White NH | 69.0 | (62.3, 75.8) | 16.0 | (10.9, 21.0) |
| | | Black NH | 53.9 | (43.8, 64.0) | 12.3 | (6.1, 18.6) |
| | | Hispanic | 61.5 | (43.0, 80.0) | 22.0 | (3.2, 40.8) |
| | | Other NH | 60.8 | (43.4, 78.3) | 16.7 | (2.3, 31.1) |
| | 45–64 years old | White NH | 75.3 | (69.3, 81.4) | 13.4 | (7.7, 19.2) |
| | | Black NH | 61.1 | (52.4, 69.7) | 10.2 | (5.0, 15.5) |
| | | Hispanic | 68.4 | (51.6, 85.2) | 18.9 | (2.4, 35.4) |
| | | Other NH | 67.8 | (51.2, 84.4) | 14.1 | (0.5, 27.7) |
| | 65 years old and older | White NH | 79.9 | (73.3, 86.5) | 7.4 | (3.0, 11.8) |
| | | Black NH | 67.1 | (56.8, 77.4) | 5.5 | (1.2, 9.8) |
| | | Hispanic | 73.8 | (57.5, 90.0) | 10.8 | (0.0, 22.5) |
| | | Other NH | 73.2 | (56.7, 89.7) | 7.8 | (0.0, 17.7) |
| Male | 18–24 years old | White NH | 56.4 | (39.4, 73.3) | 7.7 | (0.0, 15.3) |
| | | Black NH | 39.9 | (23.5, 56.4) | 5.7 | (0.0, 11.6) |
| | | Hispanic | 47.8 | (24.5, 71.1) | 11.2 | (0.0, 25.3) |
| | | Other NH | 47.1 | (23.8, 70.5) | 8.1 | (0.0, 21.7) |
| | 25–44 years old | White NH | 74.0 | (66.6, 81.4) | 16.7 | (10.8, 22.6) |
| | | Black NH | 59.8 | (49.2, 70.5) | 12.9 | (5.2, 20.7) |
| | | Hispanic | 67.1 | (49.4, 84.7) | 22.9 | (3.2, 42.6) |
| | | Other NH | 66.5 | (50.0, 83.0) | 17.4 | (3.2, 31.6) |
| | 45–64 years old | White NH | 79.7 | (73.5, 85.9) | 14.1 | (9.0, 19.1) |
| | | Black NH | 66.9 | (58.2, 75.6) | 10.8 | (5.2, 16.3) |
| | | Hispanic | 73.6 | (58.2, 88.9) | 19.7 | (3.1, 36.4) |
| | | Other NH | 73.0 | (57.8, 88.2) | 14.7 | (1.5, 28.0) |
| | 65 years old and older | White NH | 83.6 | (78.8, 88.5) | 7.8 | (3.4, 12.2) |
| | | Black NH | 72.4 | (64.4, 80.3) | 5.8 | (1.2, 10.5) |
| | | Hispanic | 78.3 | (64.4, 92.3) | 11.4 | (0.0, 23.1) |
| | | Other NH | 77.9 | (64.0, 91.8) | 8.2 | (0.0, 18.2) |

Table 8. Ever E-cigarette Use and Current E-cigarette Use Estimates and 95% Confidence Intervals for Sex by Age Category by Race/Ethnicity

| Sex | Domain | | Ever e-cigarette use | | Current e-cigarette use | |
|--------|------------------------|----------------|----------------------|---------------------|-------------------------|---------------------|
| | Age category | Race/ethnicity | Estimate (%) | Confidence interval | Estimate (%) | Confidence interval |
| Female | 18–24 years old | White NH | 50.1 | (33.1, 67.1) | 29.6 | (12.7, 46.4) |
| | | Black NH | 25.6 | (8.3, 42.9) | 18.7 | (1.1, 36.3) |
| | | Hispanic | 44.1 | (18.7, 69.5) | 16.0 | (0.0, 36.6) |
| | | Other NH | 33.6 | (10.2, 57.0) | 21.6 | (0.0, 48.4) |
| | 25–44 years old | White NH | 31.6 | (23.5, 39.8) | 16.7 | (8.6, 24.8) |
| | | Black NH | 13.6 | (7.6, 19.6) | 9.8 | (5.0, 14.7) |
| | | Hispanic | 26.6 | (8.1, 45.2) | 8.3 | (0.0, 21.3) |
| | | Other NH | 18.8 | (5.6, 32.1) | 11.5 | (0.0, 26.8) |
| | 45–64 years old | White NH | 16.9 | (11.9, 21.9) | 5.7 | (2.2, 9.3) |
| | | Black NH | 6.5 | (3.2, 9.9) | 3.2 | (0.6, 5.8) |
| | | Hispanic | 13.8 | (1.6, 26.0) | 2.6 | (0.0, 8.0) |
| | | Other NH | 9.3 | (1.8, 16.9) | 3.8 | (0.0, 11.1) |
| | 65 years old and older | White NH | 3.8 | (1.4, 6.2) | 0.0 | (0.0, 0.0) |
| | | Black NH | 1.3 | (0.2, 2.5) | 0.0 | (0.0, 0.0) |
| | | Hispanic | 3.0 | (0.0, 6.5) | 0.0 | (0.0, 0.0) |
| | | Other NH | 1.9 | (0.0, 4.0) | 0.0 | (0.0, 0.0) |
| Male | 18–24 years old | White NH | 55.5 | (35.2, 75.8) | 21.8 | (5.2, 38.4) |
| | | Black NH | 29.9 | (8.2, 51.7) | 13.2 | (0.0, 30.7) |
| | | Hispanic | 49.4 | (21.3, 77.6) | 11.2 | (0.0, 27.9) |
| | | Other NH | 38.6 | (12.0, 65.2) | 15.4 | (0.0, 36.9) |
| | 25–44 years old | White NH | 36.5 | (27.5, 45.5) | 11.6 | (5.6, 17.7) |
| | | Black NH | 16.4 | (8.7, 24.1) | 6.7 | (1.2, 12.1) |
| | | Hispanic | 31.1 | (10.5, 51.6) | 5.6 | (0.0, 14.1) |
| | | Other NH | 22.4 | (7.6, 37.3) | 7.9 | (0.0, 18.0) |
| | 45–64 years old | White NH | 20.1 | (13.2, 27.0) | 3.8 | (1.4, 6.2) |
| | | Black NH | 8.0 | (3.3, 12.7) | 2.1 | (0.0, 4.5) |
| | | Hispanic | 16.5 | (2.3, 30.8) | 1.7 | (0.0, 4.8) |
| | | Other NH | 11.3 | (2.3, 20.3) | 2.5 | (0.0, 7.0) |
| | 65 years old and older | White NH | 4.7 | (1.7, 7.7) | 0.0 | (0.0, 0.0) |
| | | Black NH | 1.6 | (0.1, 3.1) | 0.0 | (0.0, 0.0) |
| | | Hispanic | 3.7 | (0.0, 8.1) | 0.0 | (0.0, 0.0) |
| | | Other NH | 2.4 | (0.0, 5.0) | 0.0 | (0.0, 0.0) |

Table 9. Secondhand Smoke Exposure at Work and In Public Estimates and 95% Confidence Intervals for Sex by Age Category by Race/Ethnicity

| Sex | Domain | | Secondhand smoke exposure at work | | Secondhand smoke exposure in public | |
|--------|------------------------|----------------|-----------------------------------|---------------------|-------------------------------------|---------------------|
| | Age category | Race/ethnicity | Estimate (%) | Confidence interval | Estimate (%) | Confidence interval |
| Female | 18–24 years old | White NH | 6.2 | (0.1, 12.3) | 20.9 | (10.1, 31.8) |
| | | Black NH | 5.3 | (0.0, 10.9) | 18.7 | (8.0, 29.5) |
| | | Hispanic | 1.6 | (0.0, 5.2) | 5.3 | (0.0, 12.0) |
| | | Other NH | 6.9 | (0.0, 19.1) | 16.3 | (0.0, 32.7) |
| | 25–44 years old | White NH | 15.0 | (9.0, 20.9) | 36.5 | (28.3, 44.7) |
| | | Black NH | 12.8 | (6.6, 19.1) | 33.4 | (24.6, 42.1) |
| | | Hispanic | 4.1 | (0.0, 13.5) | 11.0 | (2.0, 20.1) |
| | | Other NH | 16.3 | (0.0, 32.7) | 29.8 | (12.7, 46.9) |
| | 45–64 years old | White NH | 10.2 | (5.6, 14.8) | 37.3 | (30.0, 44.5) |
| | | Black NH | 8.7 | (3.5, 14.0) | 34.1 | (25.1, 43.2) |
| | | Hispanic | 2.8 | (0.0, 9.6) | 11.5 | (1.9, 21.1) |
| | | Other NH | 11.1 | (0.0, 23.6) | 30.6 | (12.7, 48.4) |
| | 65 years old and older | White NH | 1.4 | (0.0, 3.4) | 16.9 | (11.6, 22.3) |
| | | Black NH | 1.2 | (0.0, 2.9) | 15.1 | (8.9, 21.3) |
| | | Hispanic | 0.3 | (0.0, 2.5) | 4.2 | (0.0, 8.5) |
| | | Other NH | 1.6 | (0.0, 5.5) | 13.1 | (1.8, 24.4) |
| Male | 18–24 years old | White NH | 7.2 | (0.0, 14.4) | 23.2 | (11.3, 35.1) |
| | | Black NH | 6.1 | (0.0, 13.7) | 20.8 | (9.0, 32.7) |
| | | Hispanic | 1.8 | (0.0, 5.8) | 6.1 | (0.0, 13.4) |
| | | Other NH | 7.9 | (0.0, 21.4) | 18.2 | (0.3, 36.2) |
| | 25–44 years old | White NH | 16.9 | (10.1, 23.8) | 39.5 | (31.3, 47.8) |
| | | Black NH | 14.6 | (5.0, 24.2) | 36.3 | (26.9, 45.8) |
| | | Hispanic | 4.7 | (0.0, 14.6) | 12.4 | (2.6, 22.2) |
| | | Other NH | 18.4 | (0.8, 36.1) | 32.6 | (14.4, 50.8) |
| | 45–64 years old | White NH | 11.5 | (6.0, 17.1) | 40.3 | (32.1, 48.5) |
| | | Black NH | 9.9 | (2.7, 17.2) | 37.1 | (26.6, 47.6) |
| | | Hispanic | 3.2 | (0.0, 10.5) | 12.9 | (2.4, 23.5) |
| | | Other NH | 12.6 | (0.0, 26.2) | 33.4 | (14.1, 52.7) |
| | 65 years old and older | White NH | 1.6 | (0.0, 3.7) | 18.9 | (13.2, 24.5) |
| | | Black NH | 1.4 | (0.0, 3.2) | 16.8 | (9.9, 23.8) |
| | | Hispanic | 0.4 | (0.0, 2.6) | 4.7 | (0.0, 9.4) |
| | | Other NH | 1.8 | (0.0, 5.7) | 14.6 | (2.5, 26.8) |

6. Limitations

There are two concerns with the methodology we applied: bias and variance.

There are two ways the estimates can be biased. First, there could be bias in data that is input into the model. There is not much we can do about this. Second, the model might not be correct. The model estimates the effect of region, age category, race/ethnicity, sex, and education attainment across all counties. Then, these effects are applied to the distributions in a particular county. It might be the case that a particular county behaves differently after adjusting for the covariates in the model. The same issue might occur for the 32 domains composed of the cross-classification of sex, age category (18-24, 25-44, 45-64, 65 and older), and race/ethnicity (White NH, Black NH, Hispanic, other NH).

The variance of the estimates for some of the counties and other domains is large.

7. References

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Appendix A: Technical Details about the Application of the Modeling Procedure

The methodology we applied to estimate South Carolina county-level current smoking and current vaping prevalences comes from the Srebotnjak paper (Srebotnjak et al., 2010). The methodology described in the Srebotnjak paper was applied to the BRFSS data by Dwyer-Lindgren et al. to make county-level estimates of smoking rates and prevalence of physical activity and obesity (Dwyer-Lindgren et al., 2013) (Dwyer-Lindgren et al., 2014), and to estimate county-level tobacco use and exposure in South Carolina (Eberth et al., 2018).

There were some deviations in the methodology applied to the South Carolina ATS data and the methodology described in the Srebotnjak paper. The following summarizes the differences between the models we fit and the covariate model in the Srebotnjak paper. Each difference is described in detail below.

1. We did not treat county as a random effect.
2. We used a survey modeling procedure that accounted for the sampling weights and stratification; the models in the Srebotnjak paper did not.
3. We did not run separate models by sex; the models in the Srebotnjak paper did.
4. We included the individual-level characteristic educational attainment; the models in the Srebotnjak paper did not.

Differences 1 and 2

The Srebotnjak paper treated county as a random effect. Treating county as a random effect requires fitting a random effects model. Random effects models cannot use the study design (the sampling weights and the stratification). The methodology to fit a random effects model that incorporates the study design has not been developed. Not including the sampling weights and stratification in the model yields biased results.¹ We used SAS Proc SURVEYLOGISTIC, a procedure that allows the modeling of survey data but does not allow random effects.

We had data from 1,370 respondents. The original data used for the Srebotnjak paper contained more than one million observations, and the Eberth paper used data with 7,503 observations. The lack of data in this analysis resulted in an inability to include county and county-level covariates in the modeling procedure.

¹ People without a background in survey statistics will often claim that the model incorporates the study design in the covariates. This simply is not the case, especially in the case of the Srebotnjak models that only include the individual characteristics of sex, age category, and race. In the Srebotnjak models, there no accounting for the unequal probabilities of selection due to the different sampling fractions across states. Someone from California has the same influence as someone from Wyoming, even though their weight is more than an order of magnitude larger.

From a theoretical point of view, counties are not random effects—they are fixed effects. The set of South Carolina counties is not a random sample from a theoretical collection of counties.

Difference 3

The Srebotnjak methodology ran separate models for sex: one for male and one for female. The Srebotnjak paper did not provide a good explanation for why they did this. It is possible they had convergence problems when fitting their models, which would have required them to run different models for the two sexes. Convergence is a problem when running random effects models because the procedure estimates the fixed effects and then, using that result, the random effects are estimated. It then estimates the fixed effects again using the estimates for the random effects. This process iterates until a convergence criterion is met. Sometimes the convergence criterion is never met, and the model fails to converge. In contrast, we incorporated sex as a covariate in the model and ran one fixed effects model that does not use an iterative model fitting procedure, and consequently, does not have convergence problems.

Difference 4

We included the individual-level characteristic educational attainment, in contrast to the models in Srebotnjak. Education attainment is the characteristic most correlated to smoking behaviors and most predictive of smoking behaviors. Failing to include education attainment would reduce the predictive ability of the model. Our inclusion of educational attainment is an improvement over the models in the Srebotnjak paper.

Appendix B: Describe Bootstrap Variance Estimation

We used a bootstrap method to calculate precision estimates. This method is described by J.N.K. Rao et al in a 1992 publication.²

In our application of the bootstrap method the following is applied for each outcome. Let H denote the number of sampling strata. There are $H=10$ sampling strata. Let n_h , $h=1..H$ denote the number of respondents within each of the H sampling strata. Let $w_{i,h}$, $i=1..n_h$, $h=1..H$ denote the analysis weight associated with the i^{th} respondent within the h^{th} sampling stratum. Let $B=100$ be the number of Bootstrap replicates that will be created.

Creating the Bootstrap Replicates and Bootstrap Weights

Each Bootstrap replicate and associated Bootstrap replicate weight was constructed by applying the following algorithm.

- 1) Within each sampling stratum h , select a with replacement sample of size n_h-1 from among the n_h respondents in the sampling stratum.
- 2) Count the number of times respondent i in stratum h was selected, denote this count as $m_{i,h}$
- 3) Create a data set with one record for each respondent selected in step 1.
- 4) For each sampled respondent, create the bootstrap weight, denoted $w^{B_{i,h}}$ using the following formula:

$$w^{B_{i,h}} = w_{i,h} * m_{i,h} * (n_h / (n_h - 1))$$

Apply the algorithm B times to create B replicate samples and B associated Bootstrap weights.

Creating the Estimates for each Replicate Sample

- 1) For each Bootstrap sample, create the predicted rates using the Bootstrap weight and respondents selected into the Bootstrap sample.
- 2) Create the county weights by applying the weight calibration used with the full respondent sample to the predicted values created using the Bootstrap sample and Bootstrap weight.
- 3) Estimate the county rates using the calibrated weight from step 2 with the predicted values from step 1.

Calculating the Bootstrap Variance Estimator for Each County by Year

Let θ_k denote the rate estimate for a given county in a given year produced using all respondents. Let $\theta_{k,b}$, $b=1..B$, denote the rate estimate for a given county in a given year produced using replicate sample b . The variance of θ_k is estimated using the following formula:

$$var(\theta_k) = \frac{1}{B} \sum_{b=1}^B (\theta_{k,b} - \theta_k)^2$$

² Rao, J. N. K., Wu, C. F. J. and Yue, K. (1992). Some recent work on resampling methods for complex surveys, *Survey Methodology*, 18, pp.209-217.